Outline



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- Data transformations
- Mann-Whitney test

Assessing assumptions

The t-test assuming equal variances is

very sensitive to dependence,

moderately robust against unequal variance if $n_1 \approx n_2$, but much less robust if n_1 and n_2 are quite different (e.g. differ by a ratio of 3 or more).

robust against nonnormality.

Corrective actions for 2 independent samples:

Fundamental changes if problem with independence (...)

Welch t-test if σ_1 and σ_2 differ by 3-fold or more or n_1 and n_2 differ by 3-fold or more.

If non-normal distributions:

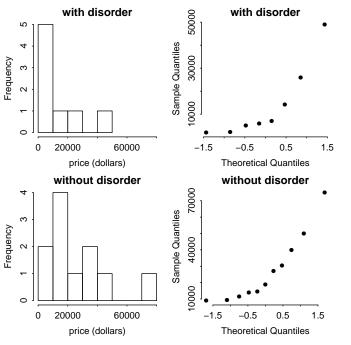
- try a data transformation,
- or switch to a non-parametric test: Mann-Whitney test.

A veterinarian wishes to know if the presence of a certain fetlock disorder in race horses affects their selling price at auction. Data on 8 horses that have the disorder, and 11 that do not (in \$)

With Disorder: 5000, 6000, 14100, 49000, 7000, 26000, 2000, 2200

Without Disorder: 27000, 14000, 11500, 19000, 9500, 40000, 75000, 9000, 14500, 50000, 30500

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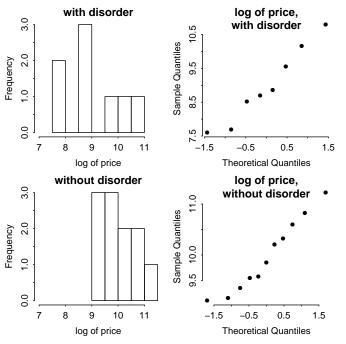
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Both samples are skewed right. Look at the log-values of the prices:

> dis
[1] 5000 6000 14100 49000 7000 26000 2000 2200
> log(dis)
[1] 8.52 8.70 9.55 10.80 8.85 10.17 7.60 7.70
> nod
[1] 27000 14000 11500 19000 9500 40000 75000 9000 14500 50000 30500
> log(nod)
[1] 10.20 9.55 9.35 9.85 9.16 10.60 11.23 9.10 9.58 10.82 10.33

Could we do the t-test on log-values instead?

If the price tends to go down with the fetlock disorder, then the log(price) also tends to be lower with the disorder than without (and vice versa).



T-test on the log-transformed prices

The distribution of log-prices looks beautifully normal for both samples! Welch t-test on the log-transformed prices:

```
dis = c(5000, 6000, 14100, 49000, 7000, 26000, 2000, 2200)
nod = c(27000, 14000, 11500, 19000, 9500, 40000, 75000, 9000,
       14500, 50000, 30500)
> t.test(log(dis), log(nod))
        Welch Two Sample t-test
data: log(dis) and log(nod)
t = -2.1955, df = 10.951, p-value = 0.05059
alternative hypothesis: true difference in means is not equal to 0
95 percent confidence interval:
-1.988749799 0.003048454
sample estimates:
mean of x mean of y
8,985856 9,978706
```

Conclusion: moderate evidence that the auction prices tend to be lower with the fetlock disorder than without (p=0.051).

Transformations

Log transformation:

helps when the distributions are skewed right, only when all values are positive

Square-root transformation:

helps when distributions are moderatetly skewed right, only when all values are ≥ 0 (zeros are okay)

Apply the same transformation (here: take the log) to all values in both samples.

Choose the transformation in order to satisfy assumptions, not based on the resulting p-value.

Confidence intervals on the original scale (not log, not transformed) are more difficult to get.

What if...

The data are too skewed and no transformation can help?

For instance: a transformation might help make one sample look normally distributed but make the other sample look worse.

Third option: use a 'non parametric' test, here test that does not assume the normal distribution: the Mann-Whitney test.

Mann-Whitney test (aka Wilcoxon rank sum test)

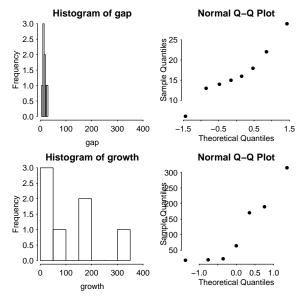
Analogous to the Wilcoxon signed-rank test (for paired samples) but here for two independent samples.

No distribution assumption, but still assume independence. Main idea: look at the ranks of the observations

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Example: Does soil respiration affect plant growth? Soil cores taken from 2 locations in a forest: under an opening in the forest canopy ("gap") and at a nearby area under heavy tree growth ("growth"). Measured: amount of carbon dioxide given off by each soil core (mol CO_2/g soil/hr). Data:

Gap 22 29 13 16 15 18 14 6 Growth 17 20 170 315 22 190 64



Gap data: distribution has normal shape,

Growth data: skewed right.

Welch t-test not recommended, but there is another way!

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```
pdf("lec15-01.pdf",width=5,height=5)
gap =c(22,29,13,16,15,18,14,6)
growth=c(17,20,170,315,22,190,64)
layout(matrix(1:4,2,2))
par(mar=c(3.1,3.1,1.5,.5), mgp=c(1.8,.4,0), tck=-0.01, las=1,bty="n")
hist(gap, xlim=c(5,400))
hist(growth, xlim=c(5,400), breaks=10)
qqnorm(gap ,pch=16)
qqnorm(growth,pch=16)
dev.off()
```

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Mann-Whitney test

*H*₀: the 2 populations have the same distribution.
 Soil respiration has the same distribution in the 2 locations, with μ₁ = μ₂ in particular.

 H_A : soil respiration does not have the same distribution in the 2 populations. Test most sensitive to a shift between the 2 distributions, so it's usually assumed that H_A is: 'the 2 distribution have different means'.

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Mann-Whitney test

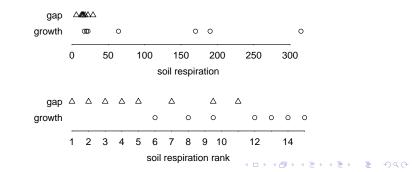
Rank the observations, calculate:

 $U_1 = \#$ of observations in group 2 that are smaller

 $U_2 = #$ of observations in group 1 that are smaller

and summarize the data by $U = \max\{U_1, U_2\}$.

If H_0 is true, then U has a Wilcoxon distribution (does not depend on the common distribution of the data).



 $\triangle \ \triangle \ \triangle \ \triangle \ \triangle$ Δ qap \triangle Δ growth 0 0 0 0 0 0 0 2 3 4 5 6 7 8 9 10 1 12 14 soil respiration rank U_1 (gap) = 0 + 0 + 0 + 0 + 0 + 1 + 2.5 + 3 = 6.5 U_2 (growth) = 5 + 6 + 6.5 + 8 + 8 + 8 + 8 = 49.5 so U = 49.5. (for ties: count 0.5) To double-check: $U_1 + U_2 = n_1 * n_2$ always. Here ves: 49.5 + 6.5 = 7 * 8 = 56

If H_0 is true: assignment of ranks to sample is completely random, so expectation: U_1 and U_2 should be similar, i.e. both intermediate, i.e. both about $n_1 * n_2/2$ (= 28 here). $U = \max\{U_1, U_2\}$ expected to be moderate.

More extreme in the direction of H_A : imbalance between U_1 and U_2 (one small, one large), i.e. large U.

Mann-Whitney test

We got U = 49.5, more extreme = larger, so p-value= $\mathbb{P}{U \ge 49.5}$.

Table E, $n_1 = 8$ and $n_2 = 7$: critical (minimum) *U* is 46 for rejecting at $\alpha = 0.05$, and 50 at $\alpha = 0.01$

So here .01 < p-value < .05

Conclusion: we have moderate evidence that the soil respiration distribution differs between the two locations.

Soil respiration has a higher mean in the aread under heavy tree growth, than in the area under the opening of the forest canopy.

Note: Table E has no number listed for $n_1 = 3$ and $n_2 = 4$: we can never reject H_0 at $\alpha = 0.05$.

One-sided Mann-Whitney test

 H_A : distribution shift with $\mu_1 > \mu_2$ for instance.

First check that the data go in the same direction as H_A , i.e. check that $U_1 > U_2$ if testing H_A : $\mu_1 > \mu_2$.

If not: p-value > 0.50.

If so: p-value is half as much as what it would be for a two-sided test.

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wilcox.test() in R

```
> gap
[1] 22 29 13 16 15 18 14 6
> growth
[1] 17 20 170 315 22 190 64
```

> wilcox.test(gap, growth)

Wilcoxon rank sum test with continuity correction

```
data: gap and growth
W = 6.5, p-value = 0.015
alternative hypothesis: true location shift is not equal
to 0
```

Warning message: In wilcox.test.default(gap, growth) : cannot compute exact p-value with ties

Warnings and Assumptions

If there are ties, the table gives approximation only.

The test does not work well if the variances are very different

To interpret H_A as simply $\mu_1 \neq \mu_2$ rather than 'the 2 distributions are different', we actually need to assume that when the 2 distributions differ, they only differ by their means –not variances.

The Mann-Whitney test is less powerful than the t-test (when both are applicable) for small sample sizes, but almost as powerful for last sample sizes.

Try transformation + t-test first: more powerful if applicable Otherwise use Mann-Whitney.