

# Review of posterior consistency & convergence rates

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# Overview of the slides

- ▶ Priors on density space
- ▶ Notions of neighborhood and distances
- ▶ Consistent tests
- ▶ Weak and strong posterior consistency - main conditions & applications
- ▶ Notion of rates of posterior convergence
- ▶ Main conditions
- ▶ Examples

- ▶ NP Bayes - priors on infinite dimensional space (density, regression function, conditional density etc)
- ▶ Examples - Dirichlet process, Gaussian process, Levy process etc
- ▶ Today - posterior consistency & rates in density estimation
- ▶  $\mathcal{X}$  - complete separable metric space ( $\mathcal{R}$  for our discussion),  $\mathcal{B}$  Borel  $\sigma$ -field on  $\mathcal{X}$
- ▶  $\mathcal{F}$  space of densities on  $(\mathcal{X}, \mathcal{B})$  w.r.t. some dominating measure
- ▶  $Y_1, \dots, Y_n \stackrel{\text{i.i.d.}}{\sim} f \in \mathcal{F}, f \sim \Pi$

- ▶ The posterior distribution is the random measure

$$\Pi(B | y^n) = \frac{\int_B \prod_{i=1}^n f(y_i) d\Pi(f)}{\int_{\mathcal{F}} \prod_{i=1}^n f(y_i) d\Pi(f)}$$

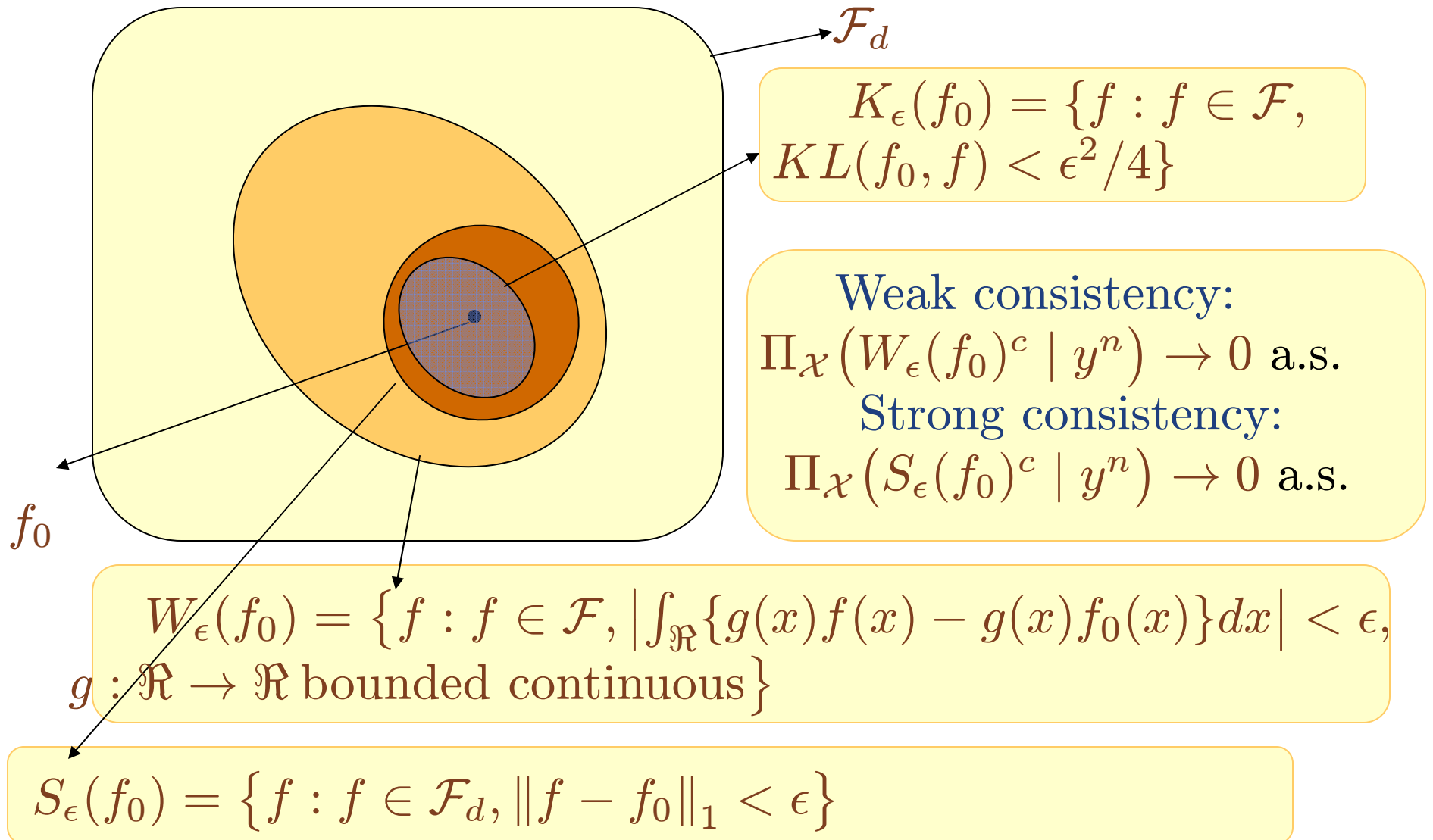
where  $B$  is a m'ble subset of  $\mathcal{F}$  and  $y^n = (y_1, \dots, y_n)$

- ▶ Assume data sampled i.i.d. from  $f_0 \in \mathcal{F}$
- ▶ Qn: does the posterior concentrate on arbitrary small neighborhoods of  $f_0$  as  $n \rightarrow \infty$ ? If so, at what rate? For which neighborhoods?
- ▶ First, need notions of distances and neighborhoods on density spaces

# Distances & nbds on density space

- ▶ Weak convergence -  $f_n \rightarrow f$  weakly if for any bounded continuous function  $\phi$ ,  $\int \phi f_n \rightarrow \int \phi f$
- ▶ A weak nbd  $W_\epsilon(f_0) = \{f \in \mathcal{F} : |\int \phi f - \int \phi f_0| < \epsilon\}$
- ▶ Strong or  $L_1$  convergence -  $f_n \rightarrow f$  in  $L_1$  if  $\int |f_n - f| \rightarrow 0$
- ▶ A strong nbd  $S_\epsilon(f_0) = \{f \in \mathcal{F} : \int |f - f_0| = \|f - f_0\|_1 < \epsilon\}$
- ▶ Also,  $\text{KL}(f_0, f) = \int f_0 \log(f_0/f)$ ,  $h^2(f, f_0) = \int (\sqrt{f} - \sqrt{f_0})^2$
- ▶ A KL nbd  $KL_\epsilon(f_0) = \{f \in \mathcal{F} : \text{KL}(f_0, f) < \epsilon\}$
- ▶ Entropy of  $\mathcal{F}_0 \subset \mathcal{F} := \log N(\epsilon, \mathcal{F}_0, \|\cdot\|_1)$  is log min. number of balls of radius  $\epsilon$  in the metric  $d$  required to cover  $\mathcal{F}_0$ .
- ▶ Interplay among these distances crucial, list of common inequalities in appendix

# Weak / strong neighborhood / consistency



- ▶ Basic idea: posterior probability of an arbitrary nbd around  $f_0$  goes to 1 as  $n \rightarrow \infty$
- ▶ Weak consistency:  $\Pi(W_\epsilon(f_0) | y^n) \rightarrow 1$  a.s.  $f_0$
- ▶ Strong consistency:  $\Pi(S_\epsilon(f_0) | y^n) \rightarrow 1$  a.s.  $f_0$
- ▶ Early result by Doob (1948): posterior consistent a.e. on prior support, not useful to check consistency at a particular density
- ▶ Breakthrough result by Schwartz (1965)

# Consistent tests

- ▶ Let  $f_0 \in \mathcal{F}$  and  $U$  be some nbd of  $f_0$
- ▶ Intuitively, should be able to separate  $f_0$  from  $U^c$  - formalized through consistent tests
- ▶ A test function  $\phi_n(y^n)$  is a non-negative measurable function bounded by  $1$
- ▶ Suppose testing  $H_0 : f = f_0$  vs  $H_1 : f \in U^c$
- ▶  $\phi_n(y^n)$  can be thought of as a randomized decision rule so that  $\phi_n(y^n) = I(\text{Rejection region} | y^n)$
- ▶ A sequence of test functions said to be **uniformly consistent** if both probabilities of type I and II errors converge to  $0$  as  $n$  increases



# Exponentially consistent & unbiased tests

- ▶  $\{\phi_n(y^n)\}$  is uniformly exponentially consistent if there exist constants  $C, \beta > 0$  such that

$$\begin{aligned} E_{f_0}[\phi_n(y^n)] &\leq C \exp(-n\beta) \\ \sup_{f \in U^c} [1 - \phi_n(y^n)] &\leq C \exp(-n\beta) \end{aligned}$$

- ▶  $\{\phi_n(y^n)\}$  is strictly unbiased if

$$E_{f_0}[\phi_n(y^n)] < \inf_{f \in U^c} [\phi_n(y^n)]$$

- ▶ The two notions above are equivalent (Hoeffding's inequality)
- ▶ Unbiased tests often easier to construct

## Theorem

Let  $\Pi$  be a prior on  $\mathcal{F}$  and  $f_0 \in \text{KL}(\Pi)$ . If there exist a sequence of exponentially consistent tests for  $H_0 : f = f_0$  vs  $H_1 : f \in U^c$ , then  $\Pi(U | y^n) \rightarrow 1$  a.s.  $P_{f_0}^\infty$

- ▶ Note  $f_0 \in \text{KL}(\Pi)$  means for any  $\epsilon > 0$ ,  $\Pi(\text{KL}_\epsilon(f_0)) > 0$
- ▶ Loosely speaking, Schwartz's theorem states large KL support + model identifiability condition  $\implies$  posterior consistency
- ▶ The KL distance related to likelihood ratios, since  $(1/n) \sum_{i=1}^n \log\{f_0(Y_i)/f(Y_i)\} \rightarrow \text{KL}(f_0, f)$  by SLLN

## Why Schwartz' s theorem works?

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Exp. Cons. sequence of tests for  $f = f_0$   
vs.  $f \in U^c$  makes likelihood ratio small

$$\Pi(U^c | \mathbf{Y}^n) = \frac{\int_{U^c} \prod_{i=1}^n \frac{f(Y_i)}{f_0(Y_i)} \Pi(df)}{\int_{\mathcal{F}} \prod_{i=1}^n \frac{f(Y_i)}{f_0(Y_i)} \Pi(df)}$$

$\longrightarrow N_n$   
 $\longrightarrow D_n$

$$\int_{\mathcal{F}} e^{-\sum_{i=1}^n \log \frac{f_0(Y_i)}{f(Y_i)}} \Pi(df) \rightarrow \int_{\mathcal{F}} e^{-nKL(f_0;f)} \Pi(df)$$

$$f_0 \in KL(\Pi) \Rightarrow \liminf e^{n\beta} D_n = \infty, \forall \beta > 0.$$

# Specialized conditions for weak and strong consistency

- ▶ Turns out that the exponentially consistent test criterion is difficult to verify
- ▶ Need easy to verifiable conditions specific to neighborhoods

Theorem: weak

If  $f_0 \in \text{KL}(\Pi)$ , the posterior is weakly consistent at  $f_0$ .

# Specialized conditions for weak and strong consistency

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## Theorem: weak

If  $f_0 \in \text{KL}(\Pi)$ , the posterior is weakly consistent at  $f_0$ .

## Theorem: strong (Ghosal et al. 1999)

If  $f_0 \in \text{KL}(\Pi)$  and there exists a sequence of subsets  $\mathcal{F}_n \subset \mathcal{F}$  such that for any  $\epsilon > 0$

1.  $\log N(\epsilon, \mathcal{F}_n, \|\cdot\|_1) \approx o(n)$
2.  $\Pi(\mathcal{F}_n^c) \leq e^{-cn}$

then the posterior is  $L_1$ -consistent at  $f_0$ .

## Weak consistency

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**Weak consistency:** If  $U_\phi$  is a weak neighborhood of  $f_0$ ,  
for a bounded conts. function  $\phi$

$$U_\phi = \left\{ f : \left| \int \phi f - \int \phi f_0 \right| < \epsilon \right\}$$

Choose the test function to be  $\phi$  since

Type I error:  $E_{f_0} \{ \phi(Y_1) \} = \int \phi f_0$  and

Power:  $\inf_{f \in U_\phi^c} \int \phi f \geq \int \phi f_0 + \epsilon$

$\Rightarrow$  existence of unbiased sequence of tests

**KL condition suffices for weak consistency**

## Strong consistency – Why Ghosal et al. 1999 works?

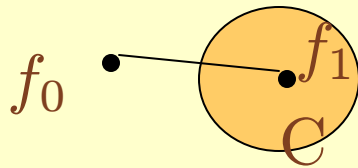
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**Strong consistency:** If  $U$  is a strong nhbd. of  $f_0$  i.e.

$$U = \{f : \|f - f_0\|_1 < \epsilon\}$$

Trivial to construct exponential consistent tests for

$$H_0 : f = f_0 \text{ \& } H_1 : f \in C$$



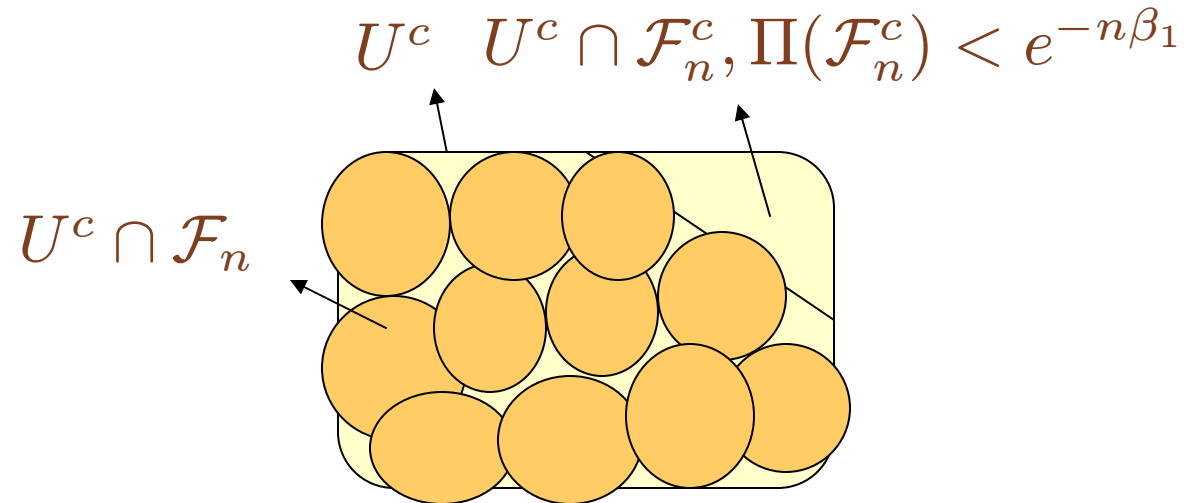
How do we do it?

$$C = \{f : \|f - f_1\|_1 \leq \|f_1 - f_0\|_1 / 2\}$$

Take  $B = \{y : f_1(y) > f_0(y)\}$  and  $\Phi = I_B$

$$\text{Then } E_{f_1}(\Phi) \geq E_{f_0}(\Phi) + \|f_1 - f_0\|_1 / 2$$

## Why Ghosal et al 1999 works?



$$N_n \leq \int_{U^c \cap \mathcal{F}_n} \prod_{i=1}^n \frac{f(Y_i)}{f_0(Y_i)} \Pi(df) + \int_{\mathcal{F}_n^c} \prod_{i=1}^n \frac{f(Y_i)}{f_0(Y_i)} \Pi(df)$$

$T_{1n}$

$T_{2n}$

$$T_{1n} \leq N(\epsilon, \|\cdot\|, \mathcal{F}_n) e^{-n\beta_{01}}, \quad T_{2n}/D_n < e^{-n\beta_{02}}$$

For  $L_1$  consistency, we need

$$\begin{aligned} \Pi(\mathcal{F}_n^c) &< e^{-n\beta_1}, \\ \log N(\epsilon, \|\cdot\|, \mathcal{F}_n) &\approx o(n) \end{aligned}$$



# Example: Density estimation using DPM

- ▶  $Y_1, Y_2, \dots, \sim f_0 \in \mathcal{F}$ , want to estimate  $f_0$
- ▶ We specify  $\Pi$  by  
 $Y_i \sim N(\mu_i, \sigma_i^2), (\mu_i, \sigma_i^2) \mid \mathcal{P} \sim \mathcal{P}, \mathcal{P} \sim DP(\alpha G_0)$ ,  $G_0$  a distribution on  $\mathcal{R} \times \mathcal{R}^+$ ,  $\pi_h$  are constructed by stick-breaking  $Beta(1, \alpha)$  variates.
- ▶ Induced density of  
 $Y_i, f(y_i) = \sum_{h=1}^{\infty} \pi_h N(y_i, \mu_h, \sigma_h^2), (\mu_h, \sigma_h)^2 \sim G_0$
- ▶ Under what conditions on  $f_0$  and  $G_0$  do we have weak and strong posterior consistency?

## Weak cons. in DPM (Ghosal et al. 1999; Tokdar 2006)

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Split  $KL(f_0, f)$  into 2 parts

$$KL(f_0, f) = \int_{\mathfrak{R}} f_0(y) \log \frac{f_0(y)}{\tilde{f}(y)} dy + \int_{\mathfrak{R}} f_0(y) \log \frac{\tilde{f}(y)}{f(y)} dy.$$

$T_1$

Impose tail conditions on  $f_0$  to approximate  $f_0$  by compactly supported  $\tilde{G}$  needed to construct density  $\tilde{f}$

$$\begin{aligned}\tilde{f} &= \int \phi\left(\frac{y-\mu}{\sigma}\right) \tilde{G}(\mu, \sigma) \\ f &= \int \phi\left(\frac{y-\mu}{\sigma}\right) G(\mu, \sigma)\end{aligned}$$

$T_2$

Verify  $\tilde{f}$  in the weak support of  $\Pi$  and also  $T_2$  is arbitrarily small

## Constructing $\tilde{f}$ : approximation idea

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Know  $\int_{\mathfrak{R}} \frac{1}{\sigma} \phi\left(\frac{y-\mu}{\sigma}\right) f_0(\mu) d\mu \rightarrow f_0(y)$  as  $\sigma \rightarrow 0$

Want compactly supported  $\tilde{G}_n$  s.t.  
 $\int_{\mathfrak{R}} \frac{1}{\sigma} \phi\left(\frac{y-\mu}{\sigma}\right) d\tilde{G}_n(\mu, \sigma) \rightarrow f_0(y)$

$d\tilde{G}_n(\mu, \sigma) \propto \delta_{\sigma_n}(\sigma) f_0(\mu) I_{[-n, n]}(\mu)$   
with  $\sigma_n \rightarrow 0$

## Handling $T_1$

A1.  $f_0$  is nowhere zero, continuous and bounded by  $M < \infty$ .

A2.  $|\int_{\mathbb{R}} f_0(y) \log f_0(y) dy| < \infty$ .

A3.  $|\int_{\mathbb{R}} f_0(y) \log \frac{f_0(y)}{\psi(y)} dy| < \infty$ ,

where  $\psi(y) = \inf_{t \in [y-1, y+1]} f_0(t)$ .

A4.  $\exists \eta > 0$  such that  $\int_{\mathcal{Y}} |y|^{2(1+\eta)} f_0(y) dy < \infty$ .

Under (A1)-(A4), using a compactly supported sequence  $\tilde{G}_n$ ,

$$f_n(y) = \int \frac{1}{\sigma} \phi\left(\frac{y - \mu}{\sigma}\right) d\tilde{G}_n(\mu, \sigma)$$

approximates  $f_0(y)$  and makes  $T_1$  arbitrarily small as  $n \rightarrow \infty$ . Choose  $\tilde{f} = f_{n_0}$  for large enough  $n_0$ .

## Handling $T_2$

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Find a weak nhbd  $W$  of  $\tilde{G}_{n_0}$  such that for  $G \in W$ ,  $T_2$  is small.

$$T_2 = \int_{\mathcal{R}} f_0(y) \log \left\{ \frac{\phi * \tilde{G}_{n_0}(y)}{\phi * G(y)} \right\} dy < \epsilon$$

$$\int_{|y| \leq k} \cdot$$
$$W = \{G : |\phi * \tilde{G}_{n_0}(y_i) - \phi * G(y_i)| < \epsilon\}$$

$$\int_{|y| \geq k} \cdot$$

use tail condition

What are the pieces left ?

Need to ensure that a DP assigns some mass at  $W$

**TRUE** if  $G_0$  has full support

# Strong consistency in DPM (Sieve construction)

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How do we construct a sieve  $\mathcal{F}_n$  such that

1.  $\log N(\mathcal{F}_n, \|\cdot\|, \epsilon) = o(n)$
2.  $\Pi(\mathcal{F}_n^c) = O(e^{-n})$

Ghosal et al. 1999 restrictive in terms of applicability

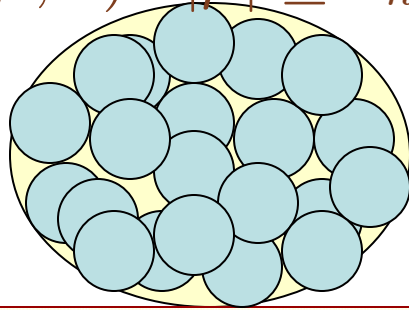
An alternative (Pati, Dunson & Tokdar, 2011)

$\mathcal{F}_n$  resembles finite mixtures  $\sum_{h=1}^{m_n} \pi_h \frac{1}{\sigma_h} \phi\left(\frac{y-\mu_h}{\sigma_h}\right)$

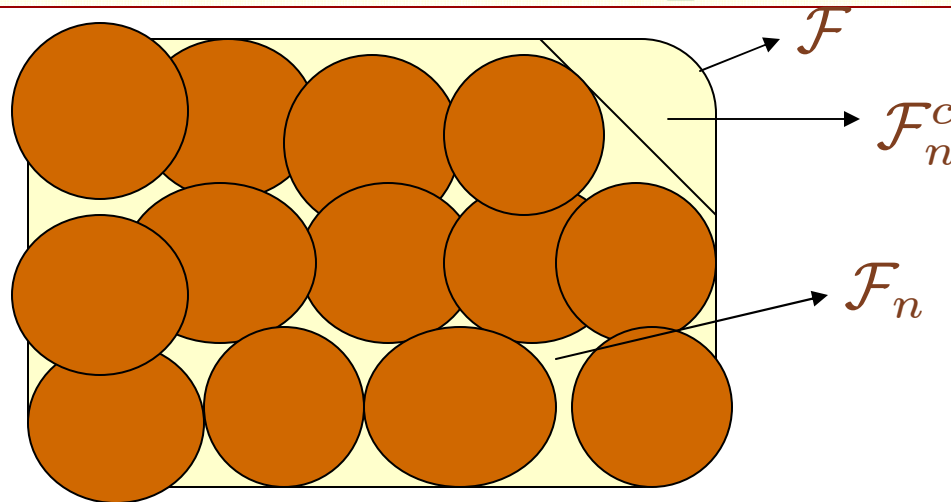
1. First few atoms are in a compact set
2. Tail sum is small

# Strong consistency in DPM (Sieve construction)

$$\Theta_{a_n, h_n, l_n} = \{(\mu, \sigma) : |\mu| \leq a_n, l_n \leq \sigma \leq h_n\}.$$



$$\mathcal{F}_n = \left\{ f : f(y) = \sum_{h=1}^{\infty} \pi_h \frac{1}{\sigma_h} \phi\left(\frac{y - \mu_h}{\sigma_h}\right), \{(\mu_h, \sigma_h)\}_{h=1}^{m_n} \right. \\ \left. \in \Theta_{a_n, h_n, l_n}, \sum_{h \geq m_n + 1} \pi_h \leq \epsilon \right\}.$$



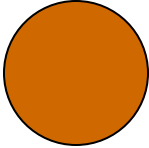
## Sieve construction (Contd.)

For  $f_1, f_2 \in \mathcal{F}_n$ ,  $\|f_1 - f_2\|_1 \leq$

$$\int_{\mathcal{X}} \sum_{h=1}^{m_n} \pi_h^{(1)} \left| \phi_{\mu_h^{(1)}, \sigma_h^{(1)}}(y) - \phi_{\mu_h^{(2)}, \sigma_h^{(2)}}(y) \right| dy$$

$$+ \sum_{h=1}^{m_n} \left| \pi_h^{(1)} - \pi_h^{(2)} \right| + 2\epsilon.$$

#  balls needed =  $N(\Theta_{a_n, h_n, l_n}, \epsilon, \|\cdot\|) \leq d_1 \left(\frac{a_n}{l_n}\right) + d_2 \log \frac{h_n}{l_n} + 1.$

#  balls needed =  $N(\mathcal{F}_n, 4\epsilon, \|\cdot\|_1) \leq$

$$\left\{ d_1 \left(\frac{a_n}{l_n}\right) + d_2 \log \frac{h_n}{l_n} + 1 \right\}^{m_n} m_n^{m_n}$$



## Strong consistency (Choice of $m_n, a_n, l_n, h_n$ )

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1. If  $G_0 = N_p(\mu; \mu_0, \Sigma_0) \times IG(\sigma^2; a, b)$ , then  
 $a_n = O(\sqrt{n}), l_n = O(\frac{1}{\sqrt{n}}), h_n = e^n$ .
2.  $P(\sum_{h=m_n+1}^{\infty} \pi_h > \epsilon) \leq e^{-m_n \log m_n}, m_n = O\left(\frac{n}{\log n}\right)$
3. With these choices of  $m_n, a_n, l_n, h_n$ , given any  $\xi > 0$ ,

$$\log(N(\mathcal{F}_n, 4\epsilon, \|\cdot\|_1)) = o(n),$$

4.  $\Pi(\mathcal{F}_n^c) \leq O(e^{-n})$

# Posterior convergence rates

- ▶ Once we have consistency, natural to ask whether we can characterize how fast the posterior concentrates
- ▶ In posterior consistency, we consider a fixed ball of radius  $\epsilon$  around  $f_0$
- ▶ Let the ball around  $f_0$  shrink with  $n$  as fast as possible so that it still captures most of the posterior mass
- ▶ The minimum possible such sequence  $\epsilon_n$  such that  $E\{\Pi(f : d(f, f_0) \geq M\epsilon_n \mid y^n)\} \rightarrow 0$  is called the **rate of convergence** of the posterior

Ghosal, Ghosh & van der Vaart (2000)

Suppose that for a sequence  $\epsilon_n \rightarrow 0$  with  $n\epsilon_n^2 \rightarrow \infty$ , a constant  $C > 0$  and sets  $\mathcal{F}_n \subset \mathcal{F}$ , one has

$$\log N(\epsilon_n, \mathcal{F}_n, d) \leq C_1 n \epsilon_n^2$$

$$\Pi(\mathcal{F}_n^c) \leq C_3 \exp\{-n\epsilon_n^2(C_2 + 4)\}$$

$$\Pi\left(f_{\mu,\sigma} : \int f_0 \log \frac{f_0}{f_{\mu,\sigma}} \leq \epsilon_n^2, \int f_0 \log \left(\frac{f_0}{f_{\mu,\sigma}}\right)^2 \leq \epsilon_n^2\right) \geq C_4 \exp\{-C_2 n \epsilon_n^2\}.$$

Then, for sufficiently large  $M$ ,  $\mathbf{E}\{\Pi(f : d(f, f_0) \geq M\epsilon_n \mid y^n)\} \rightarrow 0$

- ▶ A more subtle interplay, roughly requires prior to be uniformly spread over the parameter space
- ▶  $d$  usually Hellinger or  $L_1$  metric

# Application to a specific problem

- ▶ Density estimation model (Kundu & Dunson, 2011)

$$y_i = \mu(\eta_i) + \epsilon_i, \eta_i \sim U(0, 1), \\ \epsilon_i \sim N(0, \sigma^2), (i = 1, \dots, n).$$

- ▶  $f_0$  true density,  $F_0$  c.d.f. with  $\mu_0 = F_0^{-1} : (0, 1) \rightarrow \mathfrak{R}$ , induced density  $f_{\mu_0, \sigma}(y) =$

$$\int_0^1 \phi_\sigma(y - F_0^{-1}(t)) dt = \int_{a_0}^{b_0} \phi_\sigma(y - z) f_0(z) dz$$

- ▶  $f_{\mu_0, \sigma}(y) = \phi_\sigma * f_0(y)$ , smoothness assumptions on  $f_0$  imply  $d(f_0, f_{\mu_0, \sigma}) \rightarrow 0$  as  $\sigma \rightarrow 0$
- ▶  $f_0$  compactly supported implies  $\mu_0 : [0, 1] \rightarrow [a_0, b_0]$
- ▶  $f_0$  supported on  $\mathfrak{R}$  implies  $|\mu_0(t)| \rightarrow \infty$  as  $t \rightarrow 0/1$

# Prior specification

- ▶ Prior for  $(\mu, \sigma) \in C([0, 1]) \otimes (0, \infty)$  induces a prior on the space of densities on  $(\mathcal{R}, \mathcal{B})$
- ▶ Intuition:  $\Pi_\mu$  concentrating around  $\mu_0$  and  $\Pi_\sigma$  around zero would imply  $f_{\mu, \sigma}$  places +ve probability to arbitrary nbds of  $f_0$
- ▶ Induced measure  $\nu_\mu(B) = \tilde{\lambda}(\mu^{-1}(B))$ ,  $\mu : ([0, 1], \tilde{\lambda}) \rightarrow (\mathcal{R}, \mathcal{B})$  m'ble,  $\tilde{\lambda}$  Leb. meas. on  $[0, 1]$
- ▶ Marginalizing out  $\eta_i$ , induced density  $f_{\mu, \sigma}$ ,

$$f_{\mu, \sigma}(y) = \int_0^1 \phi_\sigma(y - \mu(t)) dt = \int \phi_\sigma(y - z) \nu_\mu(dz)$$

# Review of Gaussian processes

- ▶ Want mechanism to produce random (continuous) functions.
- ▶ A random vector  $X : (\Omega, \mathcal{E}, P) \rightarrow \mathbb{R}^k$  is Gaussian if  $a^T X$  is Gaussian for any  $a \in \mathbb{R}^k$
- ▶ Let  $X : (\Omega, \mathcal{E}, P) \rightarrow (\mathcal{C}[0, 1], \|\cdot\|_\infty)$  be measurable
- ▶  $X$  is called Gaussian if  $L(X)$  is Gaussian for any linear functional  $L$
- ▶ For example,  $L(f) = f(1/2)$ ,  $L(f) = 2f(1/3) - f(3/4)$ , ...
- ▶ Clearly, for any  $(t_1, \dots, t_m)$ ,  $\sum_{i=1}^m a_i X(t_i)$  is Gaussian for any  $a \in \mathbb{R}^m$
- ▶  $(X_{t_1}, \dots, X_{t_m})$  is MVN

# Covariance kernel approach

- ▶ Specify a joint Gaussian for  $(X_{t_1}, \dots, X_{t_m})$  consistently
- ▶ Let  $C(t, s)$  be a positive definite covariance kernel, i.e.,  $\mathbf{C} = (C(t_i, t_j))$  is positive definite for any  $t_1, \dots, t_m$
- ▶  $(X_{t_1}, \dots, X_{t_m}) \sim N(0, \mathbf{C})$ , so that  $C(s, t) = \text{cov}(X_s, X_t)$
- ▶ Common examples:  $C(t, s) = \min(t, s)$ ,  
 $C(t, s) = \exp(-\kappa|t - s|)$ ,  $C(t, s) = \exp(-\kappa|t - s|^2)$  etc

# Series expansion approach

- ▶ Mercer's theorem: There exists a sequence of eigenvalues  $\lambda_h \downarrow 0$  and an orthonormal system of eigenfunctions  $\phi_h$ , such that

$$C(s, t) = \sum_{h=1}^{\infty} \lambda_h \phi_h(s) \phi_h(t)$$

- ▶ Define  $\tilde{X}(t) = \sum_{h=1}^{\infty} \lambda_h^{1/2} Z_h \phi_h(t)$ , where  $Z_h$  i.i.d.  $N(0, 1)$
- ▶  $\text{cov}(\tilde{X}_s, \tilde{X}_t) = \sum_{h=1}^{\infty} \lambda_h \phi_h(s) \phi_h(t) = C(s, t)$
- ▶ We can start with a series representation by choosing  $\lambda_h$  and  $\phi_h$ . Different choices lead to splines, neural networks, wavelets, etc



# RKHS of Gaussian processes

- ▶ In np Bayes, want priors to place positive probability around arbitrary neighborhoods of a large class of parameter values (large support property)
- ▶ The prior concentration plays a key role in determining the rate of posterior contraction
- ▶ The reproducing kernel Hilbert space (RKHS) of a Gaussian process determines the prior support and concentration
- ▶ Let  $X$  be a zero mean Gaussian process on  $[0, 1]$  with covariance kernel  $C(s, t) = E(X_s X_t)$
- ▶ The RKHS  $\mathbb{H}$  is the completion of the linear space

$$f(t) = \sum_{h=1}^m a_h C(s_h, t), \quad s_h \in [0, 1], \quad a_h \in \mathbb{R}.$$

- ▶ Intuitively, a space of functions that are similar to the covariance kernel in terms of smoothness

- ▶ If  $f_1(t) = C(s_1, t)$ ,  $f_2(t) = C(s_2, t)$ , define  $(f_1, f_2)_{\mathbb{H}} = C(s_1, s_2)$ . Extend linearly and continuously to whole of  $\mathbb{H}$
- ▶ Finite-dimensional case: let  $X \sim N_k(0, \Sigma)$ ,  $\Sigma$  pd. Then  $\mathbb{H} = \mathfrak{R}^k$ ,  $(x, y)_{\mathbb{H}} = x^T \Sigma^{-1} y$  and hence  $\|x\|_{\mathbb{H}}^2 = x^T \Sigma^{-1} x$ . Same RKHS norm on density contours!!
- ▶ The support of a mean zero Gaussian process is the closure of the RKHS. For many standard covariance kernels, the support equals  $\mathcal{C}[0, 1]$
- ▶ The rate of posterior contraction at a function  $f_0$  depends on

$$\phi_{f_0}(\epsilon) = \inf_{h \in \mathbb{H}: \|h - f_0\|_{\mathbb{H}} < \epsilon} \|h\|_{\mathbb{H}}^2 - \log \Pr(\|X\|_{\infty} < \epsilon)$$

# Back to the rates problem

- ▶ Ongoing work (Pati, Bhattacharya & Dunson, 2011) on posterior convergence rates in NL-LVM model
- ▶ Only focus on the compactly supported case here
- ▶ Analysis of non-compact case more involved as quantile function of a non-compact density not in  $C[0, 1]$
- ▶ Standard sieve available for GP priors (van der Vaart & van Zanten 2007 onwards) - clever application of Borel's inequality
- ▶ KL condition main hurdle

- ▶ Assume  $f_0$  twice continuously differentiable, optimal minimax rate in that case  $n^{-2/5}$
- ▶ Using a GP prior with squared exponential covariance kernel for  $\mu$  & an inverse-gamma prior for  $\sigma$ , we achieve the minimax rate up to a log-factor
- ▶ One has

$$\int f_0 \log \left( \frac{f_0}{f_{\mu,\sigma}} \right)^2 \leq h^2(f_0, f_{\mu,\sigma}) \left( 1 + \log \left\| \frac{f_0}{f_{\mu,\sigma}} \right\|_\infty \right)^2$$

- ▶ With  $\epsilon_n = n^{-2/5}(\log n)^\kappa$  and  $\sigma_n^4 = \epsilon_n^2$ ,

$$\left\{ \sigma \in [\sigma_n, \sigma_n + \sigma_n^b], \|\mu - \mu_0\|_\infty \lesssim O(\sigma_n^3) \right\} \subset \left\{ \int f_0 \log \frac{f_0}{f_{\mu,\sigma}} \lesssim \sigma_n^4, \int f_0 \log \left( \frac{f_0}{f_{\mu,\sigma}} \right)^2 \lesssim \sigma_n^4 \right\}.$$

## Appendix - list of common inequalities

- ▶  $\|p - q\|_1^2 \leq 4h^2(p, q) \leq 4\|p - q\|_1$
- ▶  $\text{KL}(p, q) \geq \|p - q\|_1^2/2$
- ▶  $\text{KL}(p, q) \leq h^2(p, q)\{1 + \log \|p/q\|_\infty\}$
- ▶  $p = \text{N}(\mu_1, \sigma_1^2), q = \text{N}(\mu_2, \sigma_2^2)$  with  $\sigma_2 > \sigma_1 > \sigma_2/2$ , then  $\|p - q\|_1 \leq (2/\pi)^{0.5}|\mu_1 - \mu_2|/\sigma_2 + 3(\sigma_2 - \sigma_1)/\sigma_1$

# Key references:

- ▶ Ghosal's research page  
*[http : //www4.stat.ncsu.edu/ sghosal/papers.html](http://www4.stat.ncsu.edu/~sghosal/papers.html)*
- ▶ van der Vaart's page  
*[http : //www.few.vu.nl/ aad/research.html](http://www.few.vu.nl/~aad/research.html)*
- ▶ van Zanten's page  
*[http : //www.win.tue.nl/ jzanten/research.html](http://www.win.tue.nl/~jzanten/research.html)*
- ▶ Key **consistency** references: Barron, Schervish & Wasserman, 1999; Ghosal, Ghosh & Ramamurthy, 1999; Tokdar, 2006; Tokdar & Ghosh, 2007
- ▶ Key **rates** references: Ghosal, Ghosh & van der Vaart, 2000; Ghosal & van der Vaart, 2001; Ghosal & van der Vaart, 2007; van der Vaart & van Zanten, 2007-2009.
- ▶ Several others not cited ! Our apologies - See references within these articles.