

1. ANALYSIS OF VARIANCE

SOURCE	DF	SS	MS
TREATMENT	3	5530	1843
ERROR	8	2679	335
TOTAL	11	8209	

2. (a) `> sheer=read.table("http://www.stat.wisc.edu/~st571-1/data/sheer.dat",
+ col.names=c("type","strength"))
> tapply(sheer$strength, sheer$type, stem, scale=2)`

```

45 | 8
46 | 38
47 | 39
48 |
49 | 9
50 | 4
51 | 7

45 | 5
46 |
46 | 6
47 |
47 | 6
48 | 34
48 |
49 | 23
49 | 8

48 | 3
49 |
50 | 0029
51 | 9
52 | 8
53 | 8

49 | 5
50 |
51 | 13
52 | 0
53 | 3
54 | 03
55 |
56 | 2

```

```

48 | 2
49 | 69
50 | 3
51 | 024
52 |
53 | 2

```

```

(b) > shear.lm=lm(strength~type,shear)
> anova(shear.lm)

```

Analysis of Variance Table

Response: strength

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
type	4	12227.9	3057.0	9.1994	3.451e-05 ***
Residuals	35	11630.5	332.3		
Total	39	23858.4			

(c) $X_{ij} = \mu_i + e_{ij}$ where $e_{ij} \sim N(0, \sigma^2)$

$i = 1, 2, \dots, 5, j = 1, 2, \dots, 8.$

X_{ij} is the j^{th} measurement in the i^{th} glue type.

μ_i is the mean measurement of the i^{th} glue type.

(d) Hypotheses $H_0 : \mu_1 = \mu_2 = \dots = \mu_k$ H_A : not all μ_i are equal

$$F = 9.1994$$

$$\text{P-value} = \Pr(F_{4,35} \geq 9.1994) = 0.000$$

(e) 99% C.I. for $\mu_1 - \frac{1}{4}(\mu_2 + \mu_3 + \mu_4 + \mu_5)$

$$\lambda = (1, -1/4, -1/4, -1/4, -1/4)$$

$$\begin{aligned} \sum_{i=1}^k \lambda_i \bar{X}_i \pm t_{(N-k, \frac{\alpha}{2})} \sqrt{\text{MSE}_{\text{Err}}} \sqrt{\sum_{i=1}^k \frac{\lambda_i^2}{n_i}} &= -23.3475 \pm t_{35, 0.005} \times 18.23 \times 0.3952847 \\ &= -23.3475 \pm 19.62785 \\ &= [-42.97535, -3.71965] \end{aligned}$$

\bar{X}_i can be obtained by

```
> tapply(shear$strength,shear$type, mean)
```

A	B	C	D	E
482.625	480.875	509.875	527.125	506.000

New experimental glues appear to have stronger dry shear strength than the “standard” glue.

- (f) $H_0 : (1/2)(\mu_2 + \mu_4) - (1/2)(\mu_3 + \mu_5) = 0$ vs $H_A : (1/2)(\mu_2 + \mu_4) - (1/2)(\mu_3 + \mu_5) \neq 0$
 $\lambda = (0, 1/2, -1/2, 1/2, -1/2)$

Test statistic

$$\begin{aligned} T &= \frac{\sum_{i=1}^k \lambda_i \bar{X}_i}{\sqrt{\text{MSErr}} \sqrt{\sum_{i=1}^k \frac{\lambda_i^2}{n_i}}} \\ &= \frac{-3.935}{18.23 \times 0.3535534} \\ &= -0.6105244 \end{aligned}$$

$$\text{P-value} = 2 \Pr(T_{35} \geq 0.6105244) = 0.5454588$$

Very weak evidence against H_0 .

3. (a)

```
> y1=c(18.5, 21.0, 23.5,26.2,22.5,24.0,21.5,24.8,20.5,24.5,19.5,22.5,21.5)
> y2=c(19.0,18.0,21.3,19.5,18.3,18.8)
> y3=c(14.7,16.1,20.2,16.7,18.1,14.7,18.3,17.8,17.2,15.8,17.5)
> y=c(y1,y2,y3)
> group=c(rep(1,13),rep(2,6),rep(3,11))
> data=data.frame(y=y,group=factor(group))
> fit=lm(y~group,data)
> anova(fit)
```

Analysis of Variance Table

Response: y

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
group	2	172.419	86.209	25.003	7.168e-07 ***
Residuals	27	93.096	3.448		
Total	29	265.515			

(b)

$$F = 25.00$$

$$\text{P-value} = \Pr(F_{2,27} \geq 25.003) = 0.000$$

Very strong evidence that RH affects the average weight of newly formed pupae.

- (c) Let μ_1, μ_2, μ_3 denote the mean weights pupae at 70%, 20%, and 1% RH, respectively.

Then 95% C.I. for $\mu_2 - \mu_1$ is

(Notice $\lambda = (-1, 1, 0)$)

$$\begin{aligned} \sum_{i=1}^k \lambda_i \bar{X}_i \pm t_{(N-k, \frac{\alpha}{2})} \sqrt{\text{MSErr}} \sqrt{\sum_{i=1}^k \frac{\lambda_i^2}{n_i}} &= 19.150 - 22.346 \pm t_{27,0.025} \times \sqrt{3.45} \times \sqrt{1/13 + 1/6} \\ &= -3.196 \pm 1.881 \\ &= [-5.077, -1.315] \end{aligned}$$

(d) Test for $H_0: \sigma_1^2 = \dots = \sigma_3^2$ Use Levene's Test:

medians: 22.5 18.9 17.2

abs deviances:

gp1(13): 4.0 1.5 1.0 3.7 0.0 1.5 1.0 2.3 2.0 2.0 3.0 0.0 1.0

gp2(8) : 0.1 0.9 2.4 0.6 0.6 0.1

gp3(11): 2.5 1.1 3.0 0.5 0.9 2.5 1.1 0.6 0.0 1.4 0.3

delete exactly one zero from gp 1(13obs) and gp3(11obs)

gp1(12): 4.0 1.5 1.0 3.7 1.5 1.0 2.3 2.0 2.0 3.0 0.0 1.0

gp2(8) : 0.1 0.9 2.4 0.6 0.6 0.1

gp3(10): 2.5 1.1 3.0 0.5 0.9 2.5 1.1 0.6 1.4 0.3

ANOVA table for the data above

SOURCE	DF	SS	MS	F	p
TREATMENT	2	5.28	2.64	2.44	0.108
ERROR	25	27.09	1.08		
TOTAL	27	32.37			

P-value is 0.108, which is not strongly against the homogeneity of variances.

(e) 99% C.I. for σ^2 :

$$\begin{aligned} \left[\frac{(N-k)\text{MSE}_{\text{err}}}{\chi^2_{(N-k, \frac{\alpha}{2})}}, \frac{(N-k)\text{MSE}_{\text{err}}}{\chi^2_{(N-k, 1-\frac{\alpha}{2})}} \right] &= \left[\frac{27 \times 3.45}{\chi^2_{(27, 0.005)}}, \frac{27 \times 3.45}{\chi^2_{(27, 0.995)}} \right] \\ &= [93.15/49.64, 93.15/11.81] \\ &= [1.877, 7.887] \end{aligned}$$

```
4. (a) > y=c(12.4,15.7,10.1,11.7,12.8,
+ 14.5,11.4,18.9,16.3,13.8,
+ 14.2,18.0,20.5,15.1,18.5,
+ 22.1,17.4,20.6,19.0,19.1,
+ 26.1,19.7,23.1,21.0,24.2,
+ 22.8,27.1,26.2,19.8,25.3)
> n=rep(5,6)
> group=rep(1:6,n)
> data=data.frame(weight=y,group=factor(group))
> fit=lm(weight~group,data)
> anova(fit)
```

Analysis of Variance Table

Response: weight

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
group	5	511.60	102.32	16.559	4.248e-07 ***
Residuals	24	148.30	6.18		
Total	29	659.90			

$$F = 16.56$$

$$\text{P-value} = \Pr(F_{5,24} \geq 16.56) = 0.000$$

Very strong evidence against the claim that the amount of nitrogen does not affect the weight of corn.

(b) From the plot it appears that the relationship between weight and units of N follow a linear trend: weight increases as units of N increase.

(c) Test for hypotheses:

$$H_0 : -5\mu_1 - 3\mu_2 - 1\mu_3 + 1\mu_4 + 3\mu_5 + 5\mu_6 = 0 \quad \text{vs} \quad H_A : -5\mu_1 - 3\mu_2 - 1\mu_3 + 1\mu_4 + 3\mu_5 + 5\mu_6 \neq 0$$

$$\lambda = (-5, -3, -1, 1, 3, 5)$$

Test statistic

$$\begin{aligned} T &= \frac{\sum_{i=1}^k \lambda_i \bar{X}_i}{\sqrt{\text{MSErr}} \sqrt{\sum_{i=1}^k \frac{\lambda_i^2}{n_i}}} \\ &= \frac{84.4}{\sqrt{6.18} \times 3.741657} \\ &= 9.073696 \end{aligned}$$

$$\text{P-value} = 2 \Pr(T_{24} \geq 9.073696) = 0.0000$$

Very strong evidence against H_0 , indicating the linear trend is present.

```
5. a. > cocoon=read.table("http://www.stat.wisc.edu/~st571-1/data/cocoon.dat",
+
+           col.names=c("race","length"))
> cocoon.lm=lm(length~race,cocoon)
> anova(cocoon.lm)
```

Analysis of Variance Table

Response: length

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
race	6	288.571	48.095	10.569	4.327e-08 ***
Residuals	63	286.700	4.551		
Total	69	575.271			

$k = 7, n_i = n = 10, N = 70$

$H_0 : \mu_A = \mu_B = \mu_C = \mu_D = \mu_E = \mu_F = \mu_G$

H_A : not all the means are equal

With $F = \frac{MSTrt}{MSErr} = \frac{48.095}{4.551} = 10.569$, we get $P - \text{value} = P(F_{6,63} > 10.569) = .000$. There's very strong evidence that there's some difference among the seven races.

b. Assumptions: independent samples from seven normal population with equal variances.

To verify the normality, we can make stem-and-leaf plots or normal score plots for **each** group(seperately).

c. LSD method: $LSD = t_{N-k, \alpha/2} s_p \sqrt{\frac{2}{n}} = 1.9983(2.133) \sqrt{\frac{2}{10}} = 1.906$

`sort(tapply(cocoon$length, cocoon$race, mean))`

Race F	Race C	Race B	Race A	Race G	Race D	Race E
29.000	29.200	31.200	31.700	33.600	33.900	34.300

Q method: $QD = Q_{k, N-k, \alpha} s_p \sqrt{\frac{1}{n}} = 4.31(2.133) \sqrt{\frac{1}{10}} = 2.91$

Race F	Race C	Race B	Race A	Race G	Race D	Race E
29.000	29.200	31.200	31.600	33.600	33.900	34.300

d. Test $H_{01} : \mu_A = \mu_B$, $H_{02} : \mu_A = \mu_C$, $H_{03} : \mu_A = \mu_D$, $H_{04} : \mu_A = \mu_E$, $H_{05} : \mu_A = \mu_F$, $H_{06} : \mu_A = \mu_G$.

Using Bonferroni method with $\alpha = .05$, $r = 6$, we can compute

$t_{N-k, \frac{\alpha}{2r}} s_p \sqrt{\frac{2}{n}} = 2.7241(2.133) \sqrt{\frac{2}{10}} = 2.599$.

Race F	Race C	Race B	Race A	Race G	Race D	Race E
29.000	29.200	31.200	31.600	33.600	33.900	34.300
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Only race E and F are significantly different from race A.